

Math 191 Notes, 2003 October 14

Matching Problem, Review of specific case

Matching problem, general case with n letters. Let the probability that the $r = 0$ (number of correct letters is zero).

$$PP(r = 0) = \frac{1}{n!} \sum_{\pi} (1 - I_1) \cdots (1 - I_n)$$

For any n , the probability that any specific set of n letters end up in their correct envelopes is the same as the probability for any other set of n letters.

$$PP(r = 0) = \frac{1}{n!} \sum_s (-1)^s \binom{n}{s} (n-s)! I_1 \cdots I_s$$

Thus we get that general formula from last time.

Matching Problem, General case

Let X be the random variable defining the number of “good” letters. What is $PP(X = r)$? Pick the r that will be correct. Then make all the the rest $(n - r)$ incorrect.

$$PP(X = r) = \frac{1}{n!} (n-r)! \sum_{s=0}^{n-r} (-1)^s \frac{1}{s!} = \frac{1}{r!} \sum_{s=0}^{n-r} \frac{(-1)^s}{s!}$$

As $n \rightarrow \infty$, this is a poisson distribution with $\lambda = 1$. This is one of many ways of getting that distribution.

“Probabilistic method”

Say we have 17 fenceposts in a circle. 5 of them are rotten. Prove there is at least one set of seven consecutive posts of which 3 are rotten.

Solution using probabilistic concepts: Let R_k be the number of rotten posts in $\{k+1, k+2, \dots, k+7 \pmod{17}\}$.

$$EE(R_k) = \sum_{i=1}^7 I_{k+i} = 7 \cdot \frac{5}{17} = \frac{35}{17} > 2$$

Where I_k is the indicator function for post k being rotten. Since the probability is greater than two, we know there must be event that has three fence posts.

Binomial \rightarrow Poisson

First we want to prove that Poisson is limiting case of Binomial distribution. Binomial distribution:

$$f(k) = \binom{n}{k} p^k q^{n-k} = \frac{n!}{(n-k)!k!} \left(\frac{p}{q}\right)^k q^n$$

Now let $p \rightarrow 0, n \rightarrow \infty$ hold $np = \lambda$.

$$f(k) = \left(\frac{np^k}{q}\right) \frac{1}{k!} \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n = \frac{\lambda^k}{k!} e^{-\lambda}$$

Aside: for any specified value of k , we can probably do what we were hoping and take the limit as $n \rightarrow \infty$. Also, here's a differentiation trick.

$$\int_0^{\infty} e^{-kx} dx = \frac{1}{k}$$

$$\begin{aligned} \int_0^{\infty} x^2 e^{-kx} dx &= \int_0^{\infty} \frac{d^2}{dk^2} e^{-kx} dx \\ &= \frac{d^2}{dk^2} \left(\frac{1}{k}\right) \\ &= \frac{2}{k^3} \end{aligned}$$

Expectation for binomial distribution

$$\begin{aligned} EE(k) &= \sum_{k=0}^n k f(k) = \sum k \binom{n}{k} p^k q^{n-k} \\ &= \sum_{k=0}^n p \frac{\partial}{\partial p} \left[\binom{n}{k} p^k q^{n-k} \right] \\ &= p \frac{\partial}{\partial p} \sum \binom{n}{k} p^k q^{n-k} \\ &= p \frac{\partial}{\partial p} (p+q)^n \\ &= pn(p+q)^{n-1} = np \end{aligned}$$

And take the limit to get Poisson $np = \lambda$.

Variance for binomial distribution

$$var(X) = EE(X^2) - (E(X))^2$$

$$\begin{aligned}
EE(X^2) &= \sum_{k=0}^n k^2 f(k) \\
&= \sum_{k=0}^n k^2 \binom{n}{k} p^k q^{n-k} \\
&= p \partial / \partial p \left(p \partial / \partial p \sum_{k=0}^n \binom{n}{k} p^k q^{n-k} \right) \\
&= p \partial / \partial p p \partial / \partial p (p+q)^n \\
&= p \partial / \partial p (pn(p+q)^{n-1}) \\
&= p [n(p+q)^{n-1} + pn(n-1)(p+q)^{n-2}] \\
&= np + n(n-1)p^2
\end{aligned}$$

$$var(X) = np + n^2 p^2 - np^2 - n^2 p^2 = np(1-p) = npq$$

For the Poisson, we start with binomial and $np \rightarrow \lambda, p \rightarrow 0$: $var(X) = \lambda$.

Geometric Distribution

“Geometric” distribution: How many flips k do I need to see the first head?

$$f(k) = q^{k-1} p$$

$$\sum_1^{\infty} f(k) = p \sum_{k=1}^{\infty} q^{k-1} = \frac{p}{1-q} = 1$$

Expectation:

$$EE(X) = \sum_1^{\infty} k f(k) = p \sum_{k=1}^{\infty} k q^{k-1} = p \frac{d}{dq} \left(\sum_{k=0}^{\infty} q^k \right) = p \frac{d}{dq} \frac{1}{1-q} = p \frac{1}{(1-q)^2} = \frac{p}{p^2} = \frac{1}{p}$$

Variance:

$$\begin{aligned}
\sum_{k=1}^{\infty} k^2 f(k) &= p \sum_{k=1}^{\infty} k^2 q^{k-1} = p \sum_{k=1}^{\infty} \frac{d}{dq} \left(q \frac{d}{dq} q^k \right) = p \frac{d}{dq} \left(q \frac{d}{dq} \frac{1}{1-q} \right) = p \frac{d}{dq} \left(q \frac{1}{(1-q)^2} \right) = p \left[\frac{1}{(1-q)^2} + \frac{2q}{(1-q)^3} \right] \\
var(X) &= \frac{1}{p} + \frac{2(1-p)}{p^2} - \frac{1}{p^2} = \frac{1}{p^2} - \frac{1}{p} = \frac{1-p}{p^2} = \frac{q}{p^2}
\end{aligned}$$

Negative Binomial Distribution

How many times does it take to get r 6's when rolling a die?